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EMH vs. Phenomenological models

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Summary

- Efficient market hypothesis (EMH)
 - Rational bubbles
 - Limits and alternatives
- Phenomenological models
 - Continuous-time random walks
 - Relevance for *microscopic* market models
- Conclusion and discussion

Efficient market hypothesis (I)

Roughly speaking a market is (*informationally*) efficient if all available information is optimally used used to determine asset prices at each point in time. Assuming *risk neutral* investors:

$$E[r_A(t, t+1) | I(t)] = r_F(t, t+1)$$

A: risky asset; F: risk free asset; r_A : return rate of the risky asset; r_F : return rate of the risk free asset; $I(t)$: information available to investors at time t . In terms of price $S_A(t)$ and dividends $D(t, t+1)$ payed over the period $[t, t+1)$:

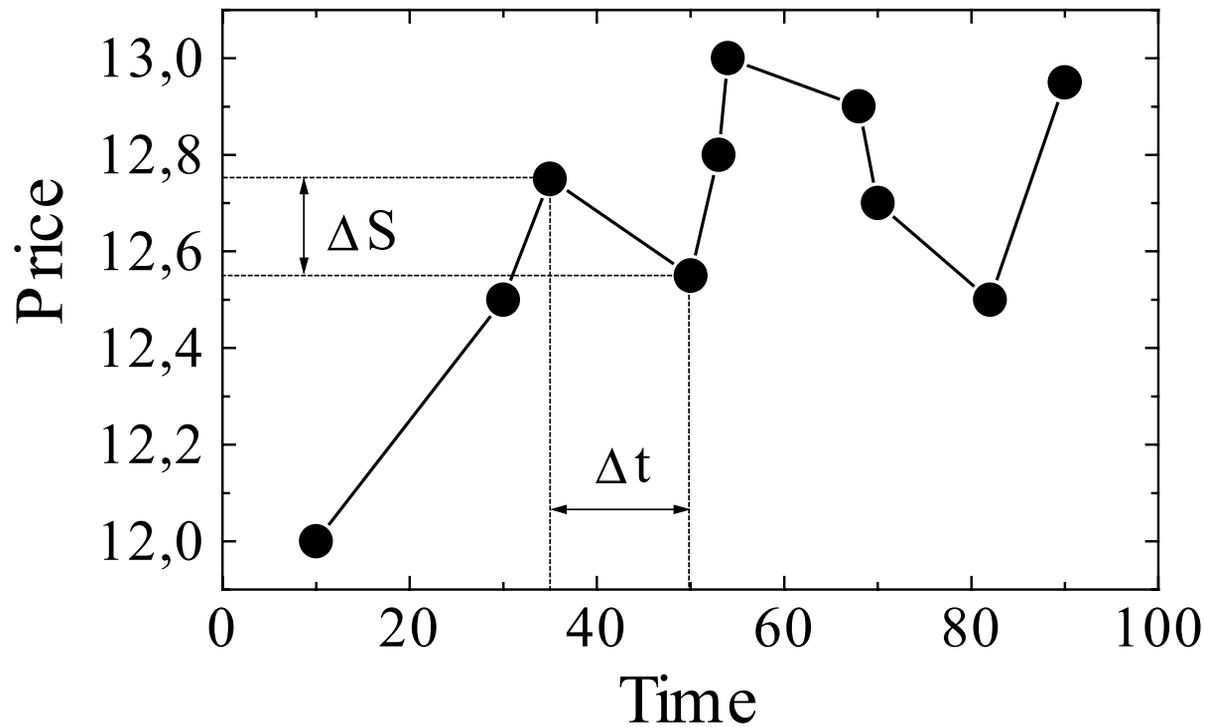
$$\frac{E[S_A(t+1) + D(t, t+1) | I(t)]}{1 + r_F(t, t+1)} = S_A(t)$$

Efficient market hypothesis (II)

- Rational bubbles do exist (bubble: price deviation from the fundamental price). The existence of bubbles is not enough to abandon the EMH. Behavioural assumptions are necessary.
- Noise traders: traders who have wrong pieces of information and believe to have all the available info.
- Zero-intelligence random traders: popular among physicists dealing with microscopic models of market dynamics.
- Our proposal: just consider market phenomenology without behavioural assumptions!

Tick-by-tick price dynamics

Price variations as a function of time



Theory (I)

Continuous-time random walk in finance (basic quantities)

$S(t)$: price of an asset at time t

$x(t) = \log[S(t)]$: log price

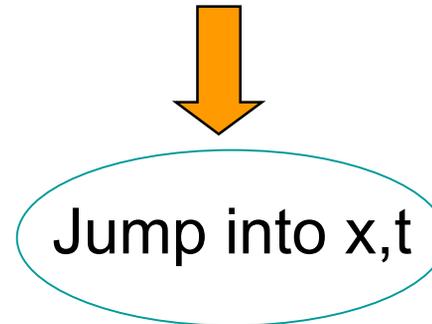
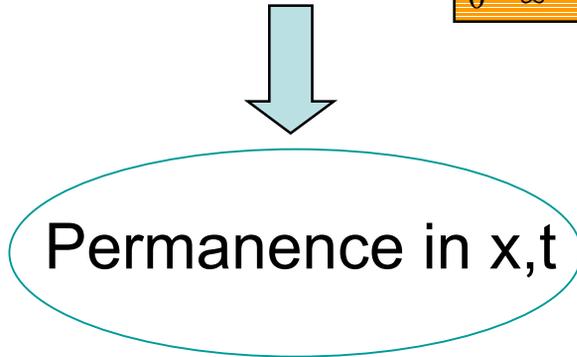
$\varphi(\xi, \tau)$: joint probability density of jumps
and of waiting times

$\xi_i = x(t_{i+1}) - x(t_i)$ $\tau_i = t_{i+1} - t_i$

$p(x, t)$: probability density function of
finding the log price x at time t

Theory (II): Master equation

$$p(x,t) = \delta(x) \Psi(t) + \int_{-\infty}^{+\infty} \int_{0}^{+\infty} \varphi(x-x', t-t') p(x', t') dt' dx'$$



$$\lambda(\xi) = \int_{0}^{+\infty} d\tau \varphi(\xi, \tau) \quad \text{Marginal jump pdf}$$

$$\psi(\tau) = \int_{-\infty}^{+\infty} d\xi \varphi(\xi, \tau) \quad \text{Marginal waiting-time pdf}$$

$$\text{In case of independence: } \varphi(\xi, \tau) = \lambda(\xi) \psi(\tau)$$

$$\Pr(\tau > \bar{\tau}) = \Psi(\bar{\tau}) = 1 - \int_{0}^{\bar{\tau}} d\tau' \psi(\tau') \quad \text{Survival probability}$$

Theory (III): Choice of marginal densities

$$\psi(\tau) = E_{\beta}(-\tau^{\beta})$$

$$E_{\beta}(z) = \sum_{k=0}^{\infty} \frac{z^k}{\Gamma(\beta k + 1)}$$

$$0 \leq \beta \leq 1$$

Mittag-Leffler function

$$\lambda(\xi) = \frac{1}{2\pi} \int_{-\infty}^{+\infty} \exp(-ik\xi) \exp(-|k|^{\alpha}) dk$$

$$0 \leq \alpha \leq 2$$

Lévy function

Results (I)

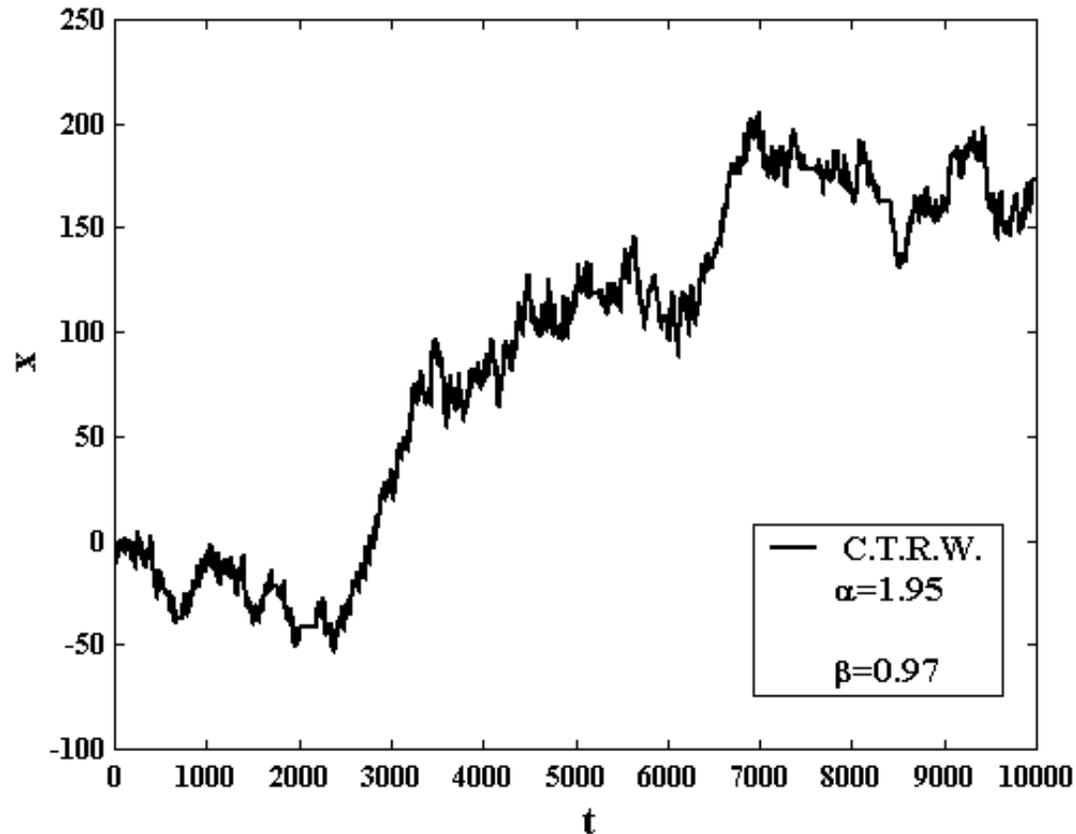


Fig. 1: Simulated log-price as a function of time. Both quantities are plotted in arbitrary units. This simulation includes nearly 8300 log-prices. It takes a few seconds to run on an old Pentium II processor at 349 MHz.

Results (II)

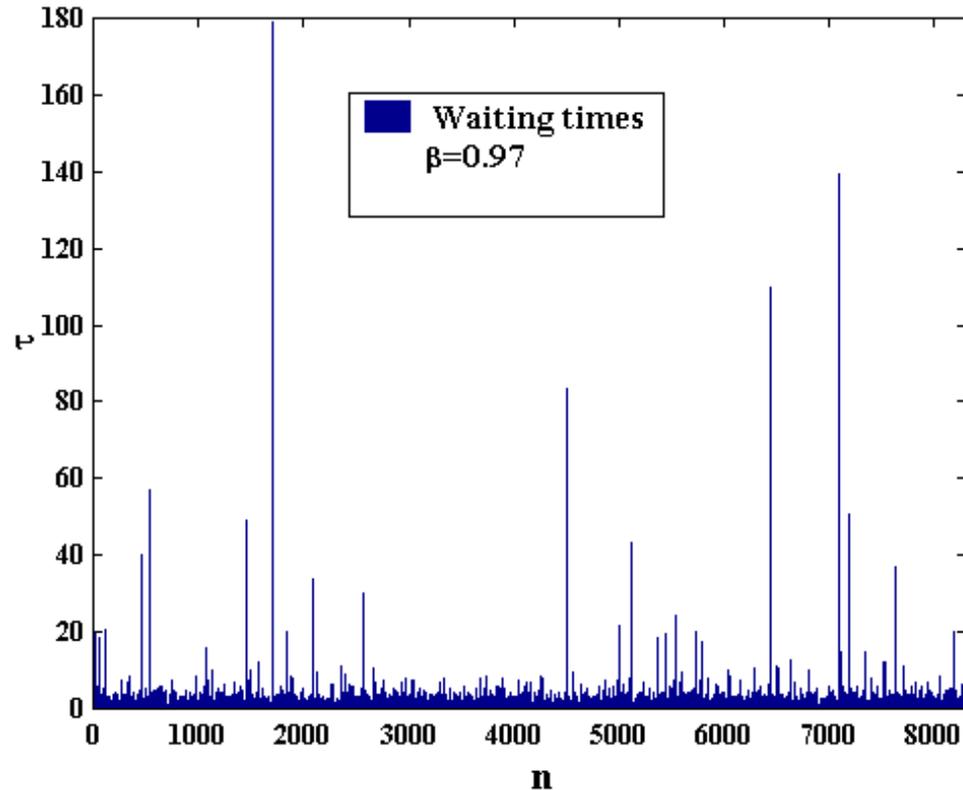


Fig 2: Waiting times as a function of the simulation index. The series has been produced by means of the rejection procedure, by comparing [0-1]-uniformly distributed deviates to a probability density of the Mittag-Leffler form with $\beta = 0.97$.

Results (III)

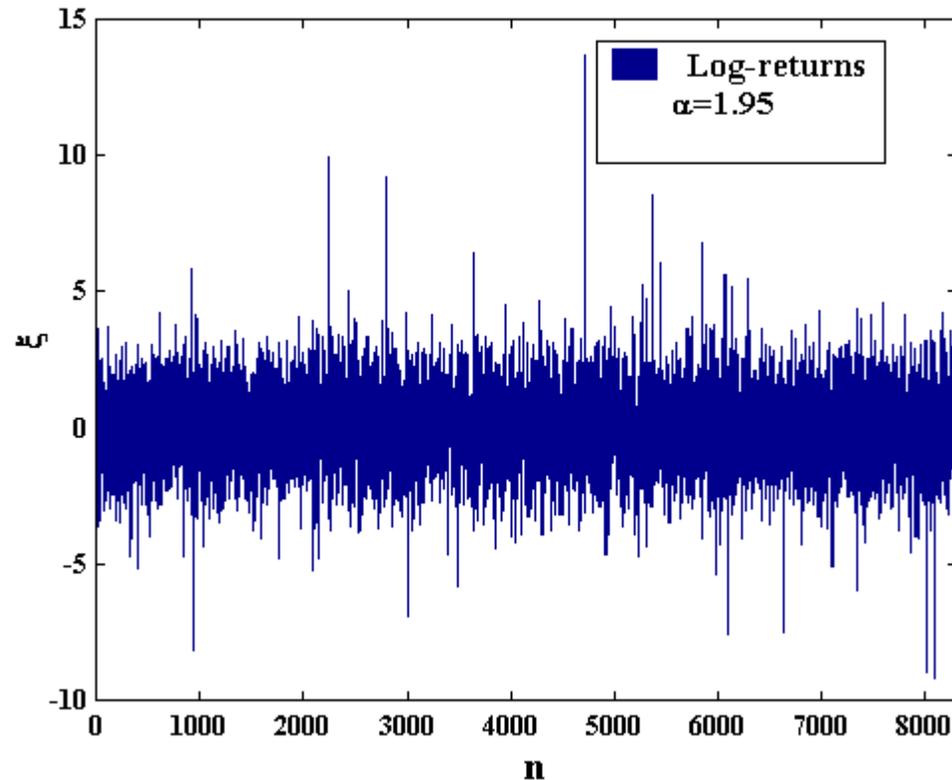


Fig. 3: Log-returns as a function of the simulation index. The series has been produced by means of the rejection procedure, by comparing [0-1]-uniformly distributed deviates with a Lévy probability density of index $\alpha = 1.95$.

Results (IV)

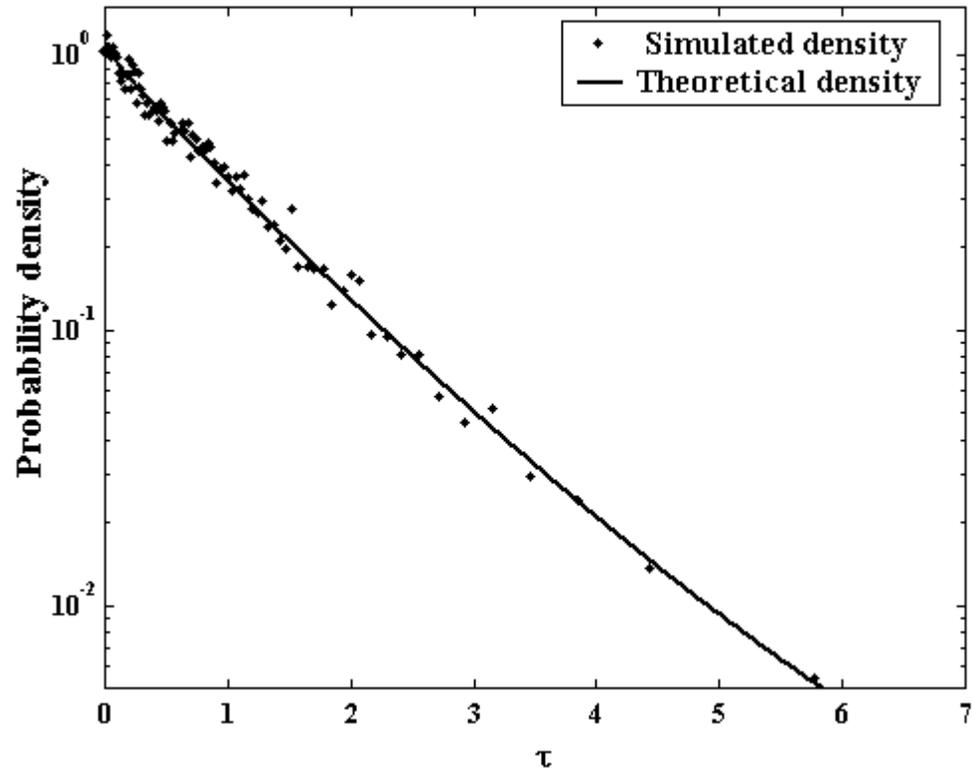


Fig. 4: Waiting times: frequency histogram of simulated data compared with theoretical values.

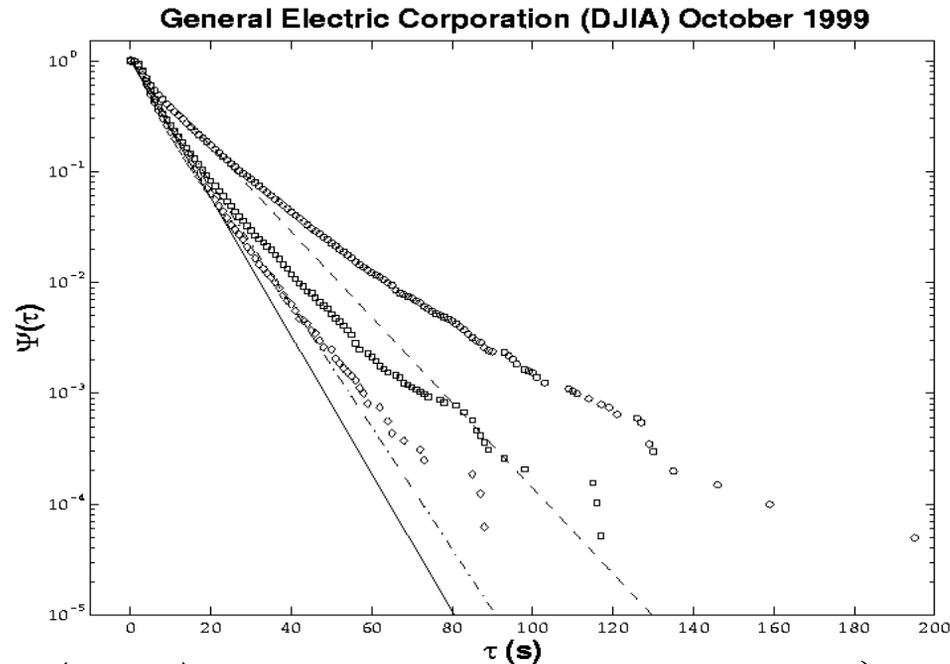
This plot is a self-consistency check for the simulation procedure.

Empirical results on the waiting-time survival function and their relevance for market models (Anderson-Darling test) (I)

Interval 1 (9-11): 16063 data; $\tau_0 = 7$ s

Interval 2 (11-14): 20214 data; $\tau_0 = 11.3$ s

Interval 3 (14-17): 19372 data; $\tau_0 = 7.9$ s



$$A^2 = \left(- \sum_{i=1}^n \frac{(2i-1)}{n} [\ln \Psi(\tau_{n+1-i}) + \ln(1 - \Psi(\tau_i))] - n \right) \times (1 + (0.6/n))$$

where $\tau_1 \leq \tau_2 \leq \dots \leq \tau_n$

$A_1^2 = 352$; $A_2^2 = 285$; $A_3^2 = 446 \gg 1.957$ (1% significance)

Empirical results on the waiting-time survival function and their relevance for market models (Anderson-Darling test) (II)

- Non-exponential waiting-time survival function now observed by many groups in many different markets (Mainardi et al. (LIFFE) Sabatelli et al. (Irish market and), K. Kim & S.-M. Yoon (Korean Future Exchange)).
- Why should we bother? This has to do both with the market price formation mechanism and with the bid-ask process. If the bid-ask process is modelled by means of a Poisson distribution (exponential survival function), its random thinning should yield another Poisson distribution. **This is not the case!**
- A clear discussion can only be found in a recent contribution by the Genoa GASM group (<http://cinef.dibe.unige.it/wehia2003/DoubleAuction.pdf>).
- Question sent to the Santa Fe group: only interlocutory answer received so far.

Results (V)

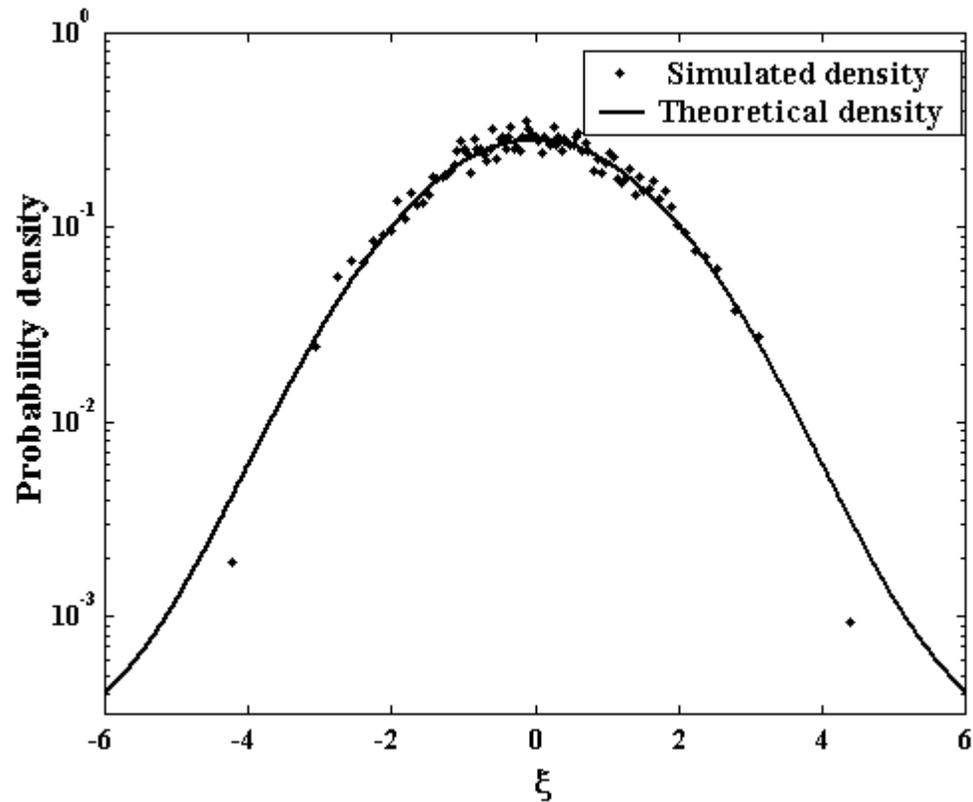


Fig 5: Log-returns: histogram of simulated data compared with theoretical values.

This plot is a self-consistency check for the simulation procedure.

Results (VI)

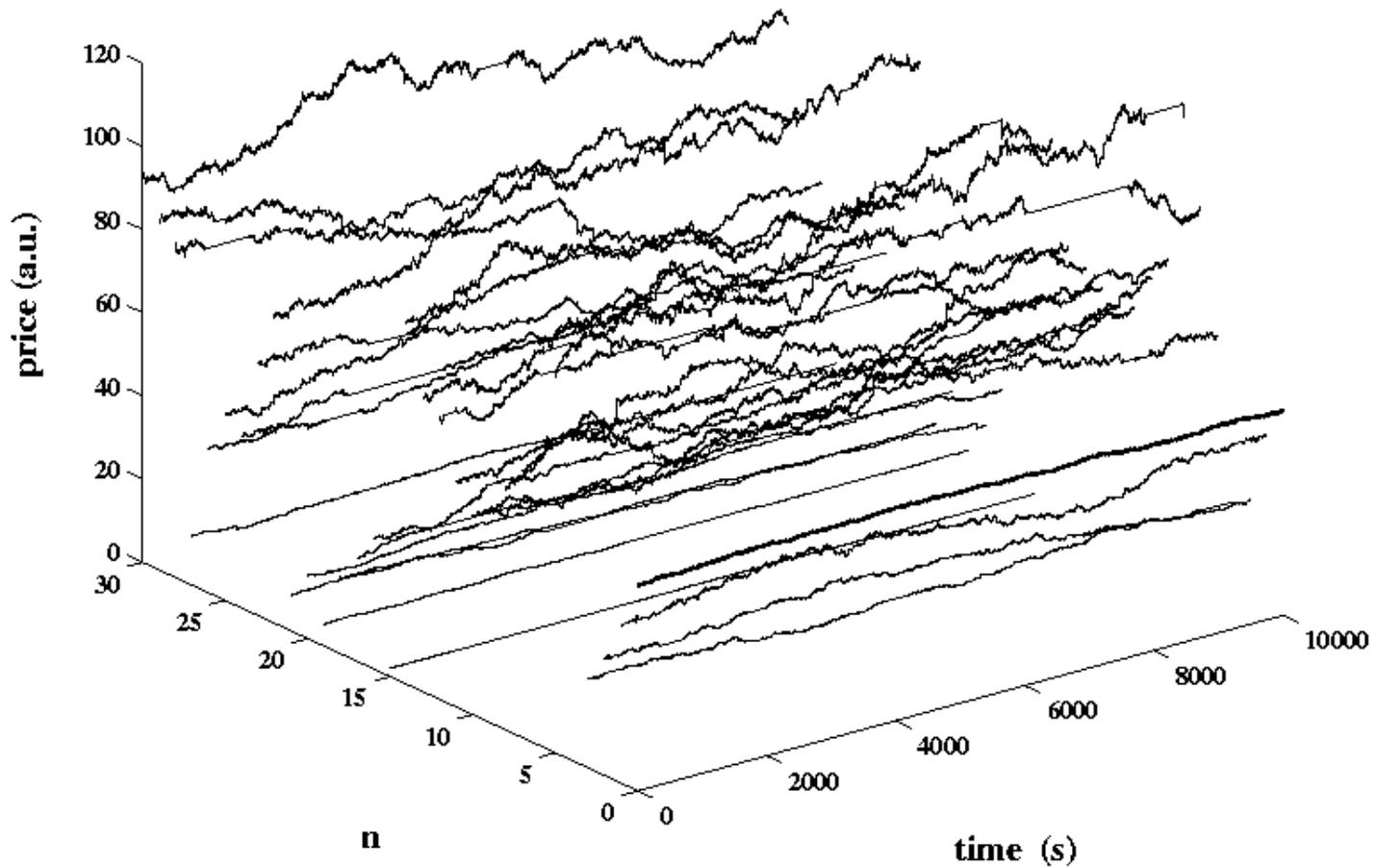


Fig 6: A synthetic market of 30 stocks.

Results (VII)

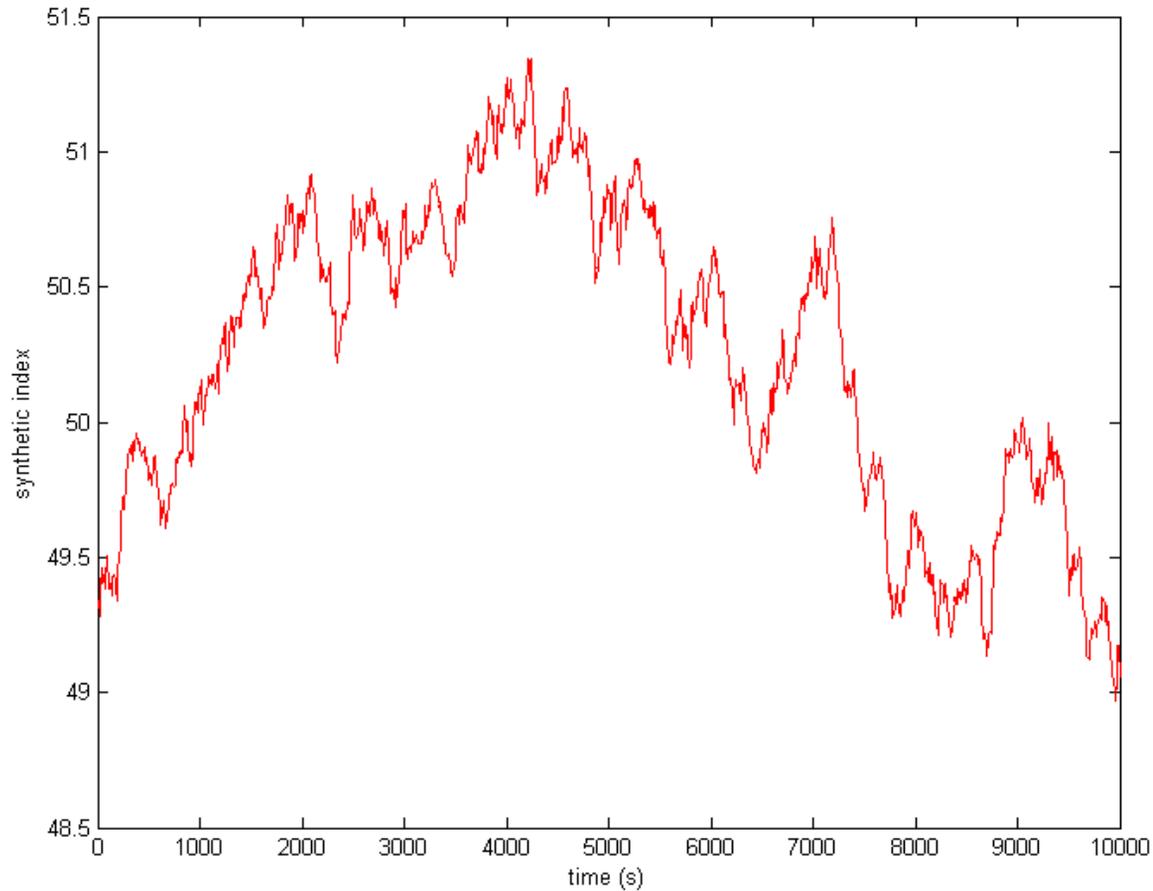


Fig 7: Synthetic market index. The index is the average of the market prices sampled every ten seconds

Conclusions

- CTRWs can be used as phenomenological models for high-frequency market dynamics.
- A synthetic market has been simulated for a specific choice of waiting-time and jump p.d.fs.
- Such simulations can be of help in various applications.

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